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Department of Mathematics



Bachelor's Thesis

The Dynamics of Vaccination Hesitancy

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Zusammenfassung

Trotz gesicherten Zusammenhangs zwischen niedriger Impfrate und dem Wiederauftauchen bereits weitgehend zurückgedrängter Infektionen, sind die durchschnittlichen Impfraten in Deutschland meist nicht hoch genug, um Herdenimmunität zu gewährleisten. Das Ziel der Bachelorarbeit ist es, dieses paradoxe Impfverhalten und die zugrundeliegenden Dynamiken durch Kombination zweier bekannter Modelle und Einbeziehung von Prozessen der Meinungsbildung zu verstehen und zu analysieren. Zunächst werden bekannte Ansätze zur Modellierung von Impfskepsis sowie Meinungsbildung mit und ohne Rücksicht auf Filterblasen präsentiert. Anschließend wird ein kombiniertes Modell, bestehend aus einem Verhalten-Inzidenz-Modell und Zelotenmodell mit Verstärkung sowie dem Konzept von Pay-off Funktionen aus der Wirtschaftsmathematik, aufgestellt und eine Bifurkationsanalyse durchgeführt. Es treten sowohl Pitchfork- als auch Sattel-Knoten-Bifurkationen für einen Spezialfall (Parameter a = 0) und eine Hopf-Bifurkation im allgemeinen Fall auf. Eine Takens-Bogdanov-Bifurkation wird vermutet, in dieser Arbeit aber nicht nachgewiesen. Im letzten Abschnitt wird das Modell auf offizielle Impfdaten gefittet, die die bundesweite Impfabdeckung gegen C-Meningokokken auf Kreisebene darstellen. Das Modell bildet die Realität besonders gut ab, was zu den Schlussfolgerungen führt, dass Kommunikation bis zu einem gewissen Grad eine größere Auswirkung auf das Impfverhalten hat als Krankheitsinzidenz. Es wird vermutet, dass die Inzidenz erst ab einem gewissen Schwellenwert größeren Einfluss auf das Verhalten hat. Außerdem arbeiten Impfgegner mit wesentlich mehr Verstärkung als Impfbefürworter, das heißt sie betonen und propagieren ihre Ansichten sehr aktiv.

Abstract

Despite the established correlation between low vaccination rates and the resurgence of infections that have already been largely suppressed, average vaccination rates in Germany are usually not high enough to guarantee herd immunity. The aim of this bachelor's thesis is to understand and analyse this paradoxical vaccination behaviour and the underlying dynamics by combining two well-known models and including processes of opinion formation. At first, known approaches to modelling vaccination hesitancy and opinion formation with and without consideration of filter bubbles are presented. Subsequently, a combined model consisting of a behaviour-incidence model and a zealot model with reinforcement as well as the concept of pay-off functions from economics will be set up and a bifurcation analysis will be performed. Pitchfork as well as saddle-node bifurcations for a special case (parameter a = 0) and a Hopf bifurcation in the general case were found. A Takens-Bogdanov bifurcation is suspected, but not proven in this work. In the last section, the model is fitted onto official vaccination data, which represent the nationwide vaccination coverage against Meningococcus C at county level in Germany. The model reflects reality pretty well, leading to the conclusion that communication has a greater impact on vaccination behaviour than disease incidence to some extent. We suppose that the influence of incidence on the behaviour increases after reaching a certain threshold. Furthermore, vaccination opponents work with much more reinforcement than vaccination proponents, i.e. they emphasise and propagate their views pretty actively.

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Abbreviations

BCM	Bounded-Confidence Model
\mathbf{BVM}	Biased Voter Model
GDR	German Democratic Republic
\mathbf{HIV}	Human Immunodeficiency Viruses
\mathbf{MMR}	Measles-Mumps-Rubella
ODE	Ordinary Differential Equation
\mathbf{SIR}	Susceptible, Infected, Recovered
STIKO	Ständige Impfkommission
WHO	World Health Organization

The notions of anti-vax, anti-immunisation and anti-vaccination as well as anti-vaxxers and vaccination opponents will be used interchangeably throughout this work.

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The R Scripts that generated Figures 3, 4, 5 and 6 can be found in Appendix 6.1 and 6.2, respectively.

1 Introduction

The personal pursuit of health, wealth and happiness may be a familiar concept to anyone. Everybody wishes to do his best on an individual base. Governmental institutions involved in Public Health also strive for health, but on a population-based scale. They aim at the best possible health conditions for the largest possible number of individuals at the same time. Interestingly, the interests and opinions of citizens and Public Health institutions do not always coincide, especially not when it comes to vaccination. As the British philosopher Jeremy Bentham saw society in the second half of the 18th century, the human was nothing more than a benefit-maximizer that pursued his personal interests without any respect of his environment [1]. So he introduced the greatest happiness principle, a concept based on the opinion that an action was considered as good if it was beneficial for the most possible individuals, and considered as bad, if it harmed the community. More or less parallel to this principle, modern medicine and research have shown the effect of herd-immunity, and how a disease can disappear within a population, if enough individuals are immunised against it [2]. Following these lines of thoughts, western governments and their Public Health institutions have shown a tendency in the past few years to impose or at least encourage vaccination [3, 4]. But this did not fit the personal needs every individual developed, valued and wanted to see pursued, so very soon, vaccine opponents appeared. One of the "anti-vax pioneers" was Reverend Edmund Massev, who lived in the 18th century and described vaccination as a "diabolical operation" [5]. The opposition to vaccination has always been present since then and has grown remarkably in the past, mostly after being empowered by studies that published allegedly shocking correlations between immunisation shots and the emergence of severe disease or disabilities [6]. A well-known example would be the study from Wakefield et al. in 1998 that pretended to have revealed the Measles-Mumps-Rubella (MMR) vaccine was very likely to predispose to autism any child who had taken the shot [7]. Despite the very small sample size n = 12, utterly speculative conclusions and even having been refuted by multiple studies [8, 9, 10], those claims received an enormous amount of publicity and

One tends to quickly classify and label individuals related to either the pro or anti-vax camp, but [11] shows that vaccination opponents as well as proponents come from various social and educational backgrounds or cultures. The main difference between those two camps seems to be the willingness to propagate their ideas amongst others. While individuals who do immunise themselves regularly are less involved in the propagation of vaccines than the government or health institutions, the so-called "anti-vaxxers" are pretty active and way less peripheral when it comes to spread their opinion [12]. Additionally, the anti-vax movement has been shown to be very influential and misleading during the opinion formation process, because of their excessive use of emotive and rhetorical appeals and their tendency to cite medical studies followed by wrongly drawn conclusions [13].

had a tremendous impact on the way people perceived vaccines [6].

Furthermore, extreme religious convictions with general rejection of vaccines may favour the reappearance of severe diseases, like the multiple outbreaks of measles in communities of Christian Scientists [14] or ultra-orthodox Jews in the state of New York [15]. In Nigeria, simple, non fact-based assumptions that the vaccine against poliovirus transmitted Human Immunodeficiency Viruses (HIV) scattered all the immunisation efforts that had been made in the previous 15 years [16]. As presented by Hussain in [17], this has not been the only case where rumours and fake news were an impactful factor for the decision finding within the population.

Apart from religious or ideological filter bubbles, vaccines are far away from being generally accepted. The results of the survey requested by the European Commission concerning Europeans' attitudes towards vaccinations [18], represented in Figure 1 below, show a non-negligible mistrust in vaccines amongst European citizens. Despite being such a minority, anti-vax groups show a remarkable impact on society and Hussain even calls it "A regression in Modern Medicine" in [17]. This work will address this phenomenon and try to understand it, by first presenting various existing models that all have tackled the problem with different creative approaches and eventually proposing an own modelling idea with the aim of integrating an opinion forming process into a compartmental modelling approach.

Figure 1:	Reasons fo	r Not	Having	Anv	Vaccination	in t	he EU	[18]
	100000110 10				1 000 01110001011	0		1-0



Base: those who did not receive a vaccination in the last five years (15,156 respondents)

2 Hesitancy Model

Firstly, we will have to take a look at a way to model vaccination hesitancy and try to understand how the individuals with a pro-vaccination opinion behave and where the dependencies of their behaviour lie. The medical/biological approach over the past few years has inter alia been to compose literature reviews and to form working groups such as in [19] to be able to define so-called "vaccine hesitancy determinants". Those are categories of reasons, for which individuals would decide not to take a vaccine. Then, a mathematical model of the dynamics had to be found, which takes into account the found determinants and one can analyse mathematically to potentially define explicit influence factors for the vaccine hesitancy dynamics.

Over time, several different models have been elaborated, but only one will be presented in this thesis: the significant work of Bauch and Bhattacharyya in [20]. They developed a behaviour-incidence model with the aim to describe the dynamics of the number of vaccinated subjects x depending on different factors such as vaccine risk, efficacy and the number of infectious persons. To get a better overview, all of the variables used in the model will be listed here:

- S : number of susceptible subjects in the population
- I: number of infected subjects in the population
- x : relative number of vaccinated subjects in the population
- μ : birth and death rate of the population
- ϵ : vaccine efficacy
- β : infection rate
- $\tau:$ case importation rate
- $\gamma:$ recovery rate
- κ : scale factor
- ω : vaccine penalty

The vaccine penalty ω basically describes the amount of risk a vaccination brings and κ is a measure for the response speed of the population to external influences. A high value for κ means that they change their opinion (pro-vaccine \rightleftharpoons contra-vaccine) immediately, a low κ stands for a very slow reaction. The variables S and I follow the convention from the well-studied Susceptible, Infected, Recovered (SIR) model developed by Kermack and McKendrick in 1927 (for more details, see [21]). We may now define the model:

(1)

Model 2.1 (Behaviour-Incidence Model by [20]). Let the context be defined as in (1). Then, the dynamics of the according behaviour-incidence model are:

$$\frac{dS}{dt} = \mu \left(1 - \epsilon x\right) - \mu S - \beta SI - \tau S$$

$$\frac{dI}{dt} = -\mu I + \beta SI - \gamma I + \tau S$$

$$\frac{dx}{dt} = \kappa x \left(1 - x\right) \left(I - \omega\right)$$
(2)

The modelling idea is that every person starts in the group of susceptible individuals Sand either stays in S, moves to the group of infected individuals I, or gets the vaccine and immediately jumps into the group of recovered subjects R. As the main interest of this model is the dynamics of vaccinated subjects x and once an individual gets into R, a return to S or I is impossible, it is assumed that the recovered have left the model, thus are not considered in Model 2.1. The dynamics of S are pretty straight-forward: the only way to get into S, is to get born into the population, so the only positive term is the birth rate μ . On the other hand, there are multiple ways to get ot of S: an individual might either die $(-\mu S)$, get infected by one individual of I $(-\beta SI)$, be born as a child of vaccinated subjects and get an immunisation shot at birth $(-\mu \epsilon x)$ or come from another population as an infected and immediately jump from S to I $(-\tau S)$.

I is designed in a similar way: newly infected subjects were either infected in the observed population $(+\beta SI)$ or came from another population and brought the disease with them $(+\tau S)$. Again, an individual leaves I at death $(-\mu I)$ or after recovering from the disease $(-\gamma I)$.

Finally, let us take a look at the dynamics of vaccinated subjects. Their growth factor consists of three parts: the response speed κ as described above, the amount of anti-vaxxers left that might switch sides (1-x) and a third term that we interpret as a "pay-off check". It computes the difference between the number of infected individuals I and the risk ω of taking the vaccine, thus basically depicts the danger emerging from the disease, as perceived by the population. When there is a high amount of infected individuals and the vaccine penalty is low, the difference will be positive and quite big, leading to a higher growth rate of x (more people will support vaccination and vaccine their children at birth) and vice versa. When ω and I are approximatively the same, the growth and decrease of x will turn out quite insignificant, respectively.

One may notice that the dynamics of the vaccinated subjects x as presented in Model 2.1 are quite basic and not realistic at all. They imply that the only influencing factor in the switch between pro and anti-vax is this kind of "pay-off comparison" described by the difference $I - \omega$ and attribute it an extreme power. According to this model, a tremendous amount of anti-vaxxers could suddenly turn their opinion around completely, despite their ideologies, fears and maybe also religious beliefs, whereas experience tells us the exact opposite. Even in such dramatic context, the path from being a staunch

anti-vaxxer to taking a flu shot into consideration can be tedious, even if the vaccine is allegedly safe or herd immunity for the whole population could be reached with that shot [22, 23, 15]. Additionally, exchange between those two camps or differently weighted opinions, such as rumours or impactful people are completely excluded from this model, yet highly present in the real world [16].

3 Opinion Model

The next step now is to understand how opinion formation works and what modelling approach fits best for us. A lot of research has been done in this field with almost every work having a different focus, due to the immense dimension of the subject. One of the most famous models is the one presented by DeGroot in [24], which aims for a consensus between all the involved agents. It is based on the presumption that one individual analyses all of the opinions of the agents connected to him and then changes his mind following an update rule that basically is weighted averaging of the personal and surrounding opinions. This model has then been refined by Friedkin and Johnsen in [25], who added the term of "innate opinion". By doing so, every agent was extended by an immutable vector that represented their opinion (values) as they would be if the agent was in a social vacuum, thus influenced by nobody. Another very modern and quite interesting modification has been done by Chitra and Musco in [26] by adding a new actor to the Friedkin-Johnsen dynamics: a network administrator. He solely is allowed to change the weights of the connections in the social network graph which resulted in a massive increase of polarization in their model. Chitra's and Musco's research has shown that a change of edge weight of 20% already lead to a polarization increase of 180%. This finding is especially relevant for the modern times, where many of such network administrators exist in form of recommendation algorithms [27]. Thus, many relevant information for opinion building on a subject may not even be delivered to every individual and falsify the final popular opinion, since there was no common information pool. Hegselmann and Krause presented an approach in [28] that modelled the interaction in social networks when considering bias and "confidence" during interaction: the Bounded-Confidence Model (BCM). The design is quite simple: every agent has a continuous opinion value that is uniformly distributed in the interval [0, 1] and only interacts with like-minded agents. To be more specific: if the other agent's opinion differs from the own opinion by less than a fixed ϵ , both opinions are changed following an update rule, otherwise there is no interaction at all. Vicario et al. extended the BCM to the Unbounded-Confidence Model [29]. The dynamics are the exact same as for the BCM, except that they also added an update rule for the agents that hold an opinion differing by more than the fixed ϵ . By doing so, they were able to replicate the observed rejection of divided opinions while still allowing some interaction, which is more realistic. Based on DeGroot's model and also considering bias and homophily, Dandekar, Goel and Lee proposed a generalization in [30]. They based their work on the fact that individuals tend to interact more with likeminded others (i.e. homophily) and on biased assimilation, a psychological phenomenon which is perfectly explained in [31]:

"Perceptions of new evidence are interpreted in such a way as to be assimilated into preexisting assumptions and expectations". According to [30], homophily alone cannot polarize society, biased assimilation is needed additionally. They also showed that some random-walk based recommending algorithms always polarised when used on biased individuals. Finally, we would like to mention the work done by Das, Gollapudi and Munagala in [32] where they defined the Biased Voter Model (BVM), which is fundamentally different from the famous Williams-Bjerknes model [33] for modelling tumour growth known under the same name. Its dynamics base on the flocking model, which basically represents the biased assimilation from above, and DeGroot's model. They aimed at modelling how different opinions are formed in social networks with informational influence. The term of informational influence has been studied by Asch in [34] and the findings were: if an individual is told to answer a question and given several anonymised responses that allegedly came from a peer group (they were in fact constructed by the experiment's conductor), the individual would still be likely to give up his own answer and stick to the popular opinion, despite the lack of group pressure in this setting.

3.1 Voter Model

As just presented, understanding how the process of forming an opinion behaves in a population and how it can be influenced or manipulated has been a huge centre of interest in the past and still is today. Several different approaches have been presented, all with various focuses due to research aim and trade-offs for simplicity's sake. Most of them base on the well-known voter model presented by Liggett in 1985 [35], where a physical approach is used and the work is built around particle behaviour and spins. But the model of interest to us is taken from a completely different field of research: population genetics. The Moran model, first introduced in 1958 by Moran [36], has originally been developed to describe the survival of two different species within one population, but can be fitted into our context pretty easily. The assumptions made in the model are:

- 1. Constant population size N
- 2. All individuals have the same chances to generate offspring (even dead individuals, this is called "selfing")
- 3. Constant birth/death rate μ

The model is defined as follows:

Model 3.1 (Moran Model by [36, 21]). Consider a total population size $N \in \mathbb{N}$, with X_t describing the number of individuals of species X at time t, Y_t those of species Y and $X_t + Y_t = N \ \forall t \in \mathbb{N}$. Then, the stochastic process counting the number of individuals of species X and Y at time t is described by following transition rates:

$$X_t \to X_t + 1 \text{ at rate } \mu Y_t \frac{X_t}{N}$$

$$(3)$$

$$X_t \to X_t - 1 \text{ at rate } \mu X_t \frac{1 - X_t}{N}$$

We may now easily reinterpret this model to depict the opinion of a population. If we consider X_t and Y_t as the number of supporters of opinion X and Y respectively and replace "death" by "not sharing this opinion any more" and "birth" by "adopting this opinion", we get exactly what we want. But research has shown that one species will die out on the long run. In our adopted voter model, an opinion would die out, which is absolutely not the case, thus this very simple approach does not fit our needs very well for the moment. This problem will be addressed in the next subsections.

3.2 Noisy Voter Model

The first modification of Model 3.1 we are going to present is the noisy voter model. It addresses the problem of one species/opinion dying out on the long run, which is not wanted at all in our setting. The line of thought is quite simple: instead of pretending linear rates for the two transitions $X_t \to X_t+1$ and $X_t \to X_t-1$, we consider probabilities. This model has been presented by Granovsky and Madras in [37]. It is a simple, ergodic variant of Liggett's voter model, as mentioned in Subsection 3.1 and reads as follows:

Model 3.2 (Noisy Voter Model by [37]). Consider a population of size $N \in \mathbb{N}$, segregated into two groups: X_t and Y_t , the supporters of opinion X or Y at time t, respectively, such that $N = X_t + Y_t$. An individual rethinks their opinion with rate μ and random opinion adoption probability p. Furthermore we consider following transition probabilities:

supporter of X sticks to his opinion/now supports Y: $p_{X,X}$ or rather $p_{X,Y}$ supporter of Y sticks to his opinion/now supports X: $p_{Y,Y}$ or rather $p_{Y,X}$ (4)

The two coherent probabilities add up to 1, respectively:

$$p_{X,Y} + p_{X,X} = p_{Y,X} + p_{Y,Y} = 1 \tag{5}$$

Then, the transition rates of the stochastic process described by the noisy voter model are:

$$X_t \to X_t + 1 \text{ at rate } \mu Y_t \left(p \frac{X_t}{N} + (1-p) p_{Y,X} \right)$$

$$X_t \to X_t - 1 \text{ at rate } \mu X_t \left(p \frac{Y_t}{N} + (1-p) p_{X,Y} \right)$$
(6)

The transition rates of this model are very similar to those of Model 3.1, with some refinements: the probabilities p and $p_{\circ,\bullet}$ with $\circ, \bullet \in \{X, Y\}$. p will be called here the random opinion adoption probability, since it depicts how likely it is that an observed individual will adopt the opinion of another individual that has been randomly chosen out of the population. Its counter probability 1 - p stands for the other possibility that the currently observed individual will adopt one of the two possible opinions without having to be "directly influenced" by another individual, no matter what the current situation in the population looks like (the opinion could not even be present in the population, but still be adopted after this step). The four different $p_{\circ,\bullet}$ describe transition probabilities, namely the probability that a supporter of X sticks to their opinion $(p_{X,X})$ or changes it $(p_{X,Y})$, analogously for supporters of Y. Now, the design of the transition rates may be explained in a pretty straight-forward manner: opinion X gains another supporter if a Y-supporter "gets influenced" by a X-supporter and adopts their opinion with probability p, or if a Y-supporter just randomly adopts one of the two opinions with probability 1-p and it turns out to be opinion X with probability $p_{Y,X}$. The explanation for the loss of a X-supporter works exactly analogously.

This adjustment of the Moran Model 3.1 seems very promising, as it eliminates the issue of an opinion dying out and presents more realistic rates to model changes in opinions. Furthermore, it adds the aspect of "opinion hopping" which is very desirable, due to its prominence in reality. Schmitt presented the most common factors that influence the vaccination rate in Germany [38]. According to german paediatricians, the most common reasons are missed appointments and illness at the planned date of the shot, which both are very accidental causalities and do not imply a strong opposition to vaccination in general. As per the parents, the most relevant factors are the importance for the personal development of the child to actually experience the illness and the frequency of side-effects of the vaccine. Those facts correspond to our assumption that most of the individuals do not have an immutable opinion on vaccines and tend to follow trends and news development (such as a higher or lower risk of a certain vaccine) as well as differentiate between the illnesses, thus make a new decision every time, whether they're going to vaccinate or not. In contrast, there is a non-negligible part of the population which changes its opinion very seldom or even not at all, due to religious, ideological or esoteric reasons [14, 23, 15]. Even though the group of anti-vaxxers may not be very representative for the population's opinion (in Germany for instance, radical immunisation opponents are estimated to represent 3-5% of the total population [22, 39, 40]), it should still be considered in our model for various reasons. Firstly, vaccination hesitancy has been classified as one of the top ten health threats in 2019 by the World Health Organization (WHO) [41]. Secondly, recall the mistrust in vaccines amongst European citizens shown in [18] and represented in Figure 1. This is not considered in Model 3.2, thus yet for us to improve.

3.3 Echo Chambers: Zealot Model

Our goal is to add the aspect of echo chambers [42] (also known as filter bubbles) into our model, since it adds more realism to the simulation. As stated in Subsection 3.2, there is a small, but non-negligible part of the population who stick to their opinion pretty strictly and uncompromisingly. The reasons for this behaviour can be various, but what is of interest for us is to integrate this fact into our model's dynamics. A big achievement in this field of study has been made by De Aguiar et al. [43, 44, 45]. They presented a model that operates on a network with $N + N_0 + N_1$ nodes, where all of the nodes either commit to state 0 or 1, which is encoded in a private state. Every single one of the Nnodes is able to change its opinion freely, those will be called "free nodes" in the following. In contrast, the N_0 and N_1 nodes are stuck with state 0 and 1, respectively. These two groups represent the part of the population which rethinks its opinion very seldom, or as in this case: never, as mentioned in Subsection 3.2. The scientific community agreed on a common name for agents with this kind of behaviour: zealots [46, 44, 45]. In [44], the model considers discrete time steps and selects a random free node at every step and updates its private state following a simple rule:

- 1. State stays the same with probability p
- 2. State of a random connected neighbour is adopted with probability 1 p

They also showed that this model already may simulate various different situations such as dynamics of small external magnetic fields, elections, spread of epidemics and population dynamics. Again, the electoral/voting approach can easily be reinterpreted as vaccination behaviour to fit our needs. When considering a fully connected network (which is also the case in our setting), De Aguiar et al. named the probability to have m free nodes in state 1 at time $t P_t(m)$ and were able to compute the dynamics of the next time step:

$$P_{t+1}(m) = P_t(m) \left[p + \frac{1-p}{N(N+N_0+N_1-1)} \times \left(m(m+N_1-1) + (N-m)(N+N_0-m-1) \right) \right] + P_t(m-1) \frac{1-p}{N(N+N_0+N_1-1)}(m+N_1-1)(N-m+1) + P_t(m+1) \frac{1-p}{N(N+N_0+N_1-1)}(m+1)(N+N_0-m-1).$$

$$(7)$$

Only the mid part of these dynamics will be clarified in the explanation of the transition rates (8), an exhaustive discussion may be found in [43, 44].

We may now present a definition of the zealot model as defined by the transition probabilities (7), but with the simple difference that we allow self-connection in the graph. In the population dynamics setting this would mean some kind of asexual reproduction, which might not be needed, but in our context it means "self-influencing" which is totally possible. The model reads as follows:

Model 3.3 (Zealot Model for Two Parties by [43, 44, 45]). Consider a population of size $N \in \mathbb{N}$, with $X_t, Y_t \in \{0, ..., N\}$ supporters of opinion X and Y at time t, respectively. Let $N = X_t + Y_t$. Furthermore, let N_X and $N_Y > 0$ be the number of zealots, that support opinion X and Y, respectively. An individual may reconsider his opinion with rate μ , select another individual or zealot and copy their opinion. The probability to be chosen during this reconsideration process is the same for all individuals and zealots. Then, the transition rates of the underlying stochastic process are:

$$X_t \to X_t + 1 \text{ at rate } \mu \left(N - X_t \right) \frac{X_t + N_X}{N + N_X + N_Y}$$

$$X_t \to X_t - 1 \text{ at rate } \mu X_t \frac{N - X_t + N_Y}{N + N_X + N_Y}$$
(8)

The transition rates might not be obvious after considering the transition probabilities (7), so a derivation will be done here. First, we will reinterpret all the terms for transparency's sake. We do not consider nodes any more, but individuals in a population and they are not in state 0 or 1, but have the opinion X or Y. $P_{t+1}(m)$ consists of three parts:

- 1. there were already X_t supporters of X and their final opinion didn't change
- 2. there was one supporter less and one individual's opinion switched from Y to X
- 3. there was one supporter more and one individual's opinion switched from X to Y

Since the reasoning for the two transition rates is the exact same, we will just derive the rate for $X_t \to X_t + 1$, thus only consider the mid probability of (7). It describes the transition $m-1 \to m$, which is exactly equivalent to what we want, given that we do not consider the number of supporters m any more, but the stochastic process X_t . $P_t(m-1)$ may be omitted (\dagger), as it is just relevant for the probabilistic approach, whether we fall into that case or not. Then, we modify the denominator to fit our approach with the "self-influencing" (\star). We know that the probability for an individual to overthink his opinion in the model of [43, 44] is 1 - p, so we may define $\mu \coloneqq 1 - p$ (\ddagger). After simply rearranging the leftover fractions and replacing the probability that N - m + 1 nodes are in state 0 with terms of the stochastic process (as defined in Model 3.3: $N = X_t + Y_t$), we get the wanted transition rate ρ from (8). In formulas:

$$P_t (m-1) \frac{1-p}{N(N+N_X+N_Y-1)} (m+N_X-1) (N-m+1)$$

$$\stackrel{(\dagger)}{\hookrightarrow} \frac{1-p}{N(N+N_X+N_Y-1)} (m+N_X-1) (N-m+1)$$

$$\stackrel{(\star)}{\hookrightarrow} \frac{1-p}{N(N+N_X+N_Y)} (m+N_X-1) (N-m+1)$$

$$\stackrel{(\dagger)}{\hookrightarrow} \frac{\mu}{N(N+N_X+N_Y)} (m+N_X-1) (N-m+1)$$

$$\stackrel{(\dagger)}{\hookrightarrow} \frac{\mu}{N(N+N_X+N_Y)} (m-1) + N_X$$

$$\stackrel{(\star)}{\hookrightarrow} \mu (N-X_t) \frac{X_t+N_X}{N+N_X+N_Y} =: \rho$$

This model looks very promising and almost suits our needs, but still lacks one aspect: reinforcement. This means that individuals, whose life is predominantly taking place in a filter bubble, are less likely to interact with a group of individuals who is representative for the population. Their social contacts will rather all be inside the filter bubble, thus all be like-minded and lead to the opinion to never change or even to enhance.

Research has shown that most of the populist parties and candidates in Europe, as well as in the USA can be determined quite easily, since their parameter of reinforcement turns out to be pretty high [47]. In Germany for instance, the extreme right-wing party "Alternative für Deutschland" and the extreme left-wing party "Die Linke" both show high reinforcement parameters. Additionally, the reinforcement of the extreme right-wing seems to be linked with the German federal states that used to be part of the German Democratic Republic (GDR), whilst there seem to be no geographical connections to the reinforcement of "Die Linke". During the US presidential election 2016, the republican's reinforcement factor tripled, most likely due to Trump's populist arguments [47]. Reinforcement seems to play a major role in those election processes, since it "boosted" an opinion, which did not depict the preponderant opinion in the population, to still get a favourable election result. We want this factor to be included in our approach, thus two models will be presented in Subsection 3.4 that take reinforcement into account.

3.4 Mean Field Voter Model and Zealot Model with Reinforcement

We will start this subsection off by presenting the mean field voter model that has been introduced by Sano et al. [48]. It is yet again a model built to describe election dynamics, but we have clarified the parallel between elections and vaccine opinion multiple times. The aspect of reinforcement mentioned above is not implemented in this model, but another interesting approach to simulate realistic dynamics between zealots and "normal", free individuals. In Model 3.3 as well as in Model 3.4, the zealots are considered as "outside" of the population, categorised into the groups N_X , and N_Y . Sano's approach is to include the zealots into the population, s.t. $N = X_t + Y_t + N_X + N_Y$ (for a two-party setting, as we have been following in this work). So, the population consists of two groups: the free voters, whose opinion may fluctuate and change, and the zealots, whose opinion is immutable. The decisive aspect are those two assumptions:

- 1. Zealots influence free voters, not inversely
- 2. Not every zealot actively influences free voters, but only a fraction ϕ_X and ϕ_Y , respectively

The transitions between opinions for a free voter are the same as described in Model 3.3. With a given rate μ , a voter rethinks his opinion, randomly picks another voter out of the population and copies their opinion. Additionally, this process may now be influenced by the zealots with rate ϕ_X and ϕ_Y , respectively, so that the change of opinion is favourable for the opinion they root for.

Sano assumes that only a part ϕ_X, ϕ_Y , respectively, of the zealots would influence the opinion of free nodes. An assumption we do not totally agree with. In our perception, the pro-vaccination zealots mainly consist of the government, health bodies and their presence in the media (public health programs, mandatory vaccines), thus they will very actively propagate their opinion. Therefore we consider every pro-vaccine zealot as active in our setting. On the other side, we assume that the anti-vax zealots behave quite the same and propagate their arguments and ideas very actively [13, 22]. Hence, we do not think that an "active influence rate", as it has been introduced by Sano will be needed in our model. Still, his work had to be presented here, because it addresses the issue in a very interesting innovative way. In the following, we will introduce the reinforced zealot model, which perfectly fits our needs. All of the new terms and notation will be explained in the subsequent paragraph. The model reads as follows:

Model 3.4 (Zealot Model for Two Parties with Reinforcement by [47]). Consider a population of size $N \in \mathbb{N}$, where $X_t, Y_t \in \{0, ..., N\}$ describe the amount of supporters of opinion X and Y at time t, respectively. Let $N = X_t + Y_t$. Furthermore, let N_X and $N_Y > 0$ be the number of zealots that support opinion X and Y, respectively. An individual of the population (not a zealot) may reconsider his opinion with rate μ via the same process as described in Model 3.3 and in addition of the contact rates θ_X, θ_Y . Then, the transition rates of the stochastic process are:

$$X_t \to X_t + 1 \text{ at rate } \mu \left(N - X_t\right) \frac{\theta_X (X_t + N_X)}{\theta_X (X_t + N_X) + (N - X_t + N_Y)}$$

$$(9)$$

$$X_t \to X_t - 1 \text{ at rate } \mu X_t \frac{\theta_Y (N - X_t + N_Y)}{(X_t + N_X) + \theta_Y (N - X_t + N_Y)}$$

When we take a look at (8), we notice there is just a slight difference between the transition rates stated there and the ones we are currently looking at. It consists of the two contact rates θ_X and θ_Y . They describe the probability of a Y-supporter getting in touch with a X-supporter (θ_X) and vice-versa. With this additional information, the derivation of the transition rates of (9) is pretty easy:

For $X_t \to X_t + 1$ we need a supporter of Y to firstly, reconsider his opinion (happens at rate μ) and secondly, actually change it to X. As stated in the opinion change dynamics in Model 3.3, which are also used here, this only happens if said supporter gets influenced by a supporter of X. So, we multiply the possible number of encounters with X-supporters $(X_t + N_X)$ with the probability θ_X of getting out of the "Y filter bubble", divided by all possible encounters. The rate of the transition $X_t \to X_t - 1$ is derived exactly analogously.

4 Presentation of the Combined Model

Lots of different modelling approaches to capture the essence of opinion formation in a (still very simplified) population and spectrum of opinions have been presented in Section 3. We looked at the transitions of some of them and concluded that Model 3.4 fits our assumptions about the behaviour of vaccine hesitancy best. In this section, we will present a model for the dynamics of vaccine hesitancy, whose two fundamental parts, the equation structure and the opinion formation process, will be inspired by Bauch's behaviour-incidence model (2.1) and by the reinforced zealot model (3.4). During model design, several approaches to render the opinion formation and influencing process in the model more realistic were considered. The notion of pay-off functions taken from the field of financial mathematics seemed to suit our needs best and was strongly investigated. For simplicity's sake, the aspirations to make use of an actual pay-off function have been put down and we introduced a pay-off parameter, as will be described in the following. Firstly, we will derive the transition rates for the stochastic process X_t as they are described by our model. Then, we will compute the deterministic limit to receive a system of Ordinary Differential Equations (ODEs), which we may then analyse.

4.1 Derivation of the Transition Rates

To be able to derive the transition rates, we have to state the model behaviour we wish for. The transition rates of (9) are already quite nice, but do not fit our needs entirely. We want to integrate the aspect of mutual influence between the individuals in the population and the personal "pay-off" everybody computes in their heads before making a somewhat important decision. One could come up with a very sophisticated and realistic "pay-off function" for both pro and anti-immunisation parties and integrate it into the model, thus tremendously raise the complexity. We figured that a simpler way of doing so would be to modify the two parts of the extended population that are already used $(X_t + N_X \text{ and } N - X_t + N_Y)$ so that they become a pro-vaccination and anti-vaccination part. Yet again, to reduce complexity, the factor considered as most relevant for the personal pay-off computation of an individual that is rather tending towards vaccinating, is the prevalence of the disease, described by aI with $a \in \mathbb{N}$ and I as in (1). Whereas for the anti-vaccination part, we introduce another factor $\Omega(I): [0,1] \to \mathbb{R}^+$ that describes an anti-vaccination influence in the thought process (an anti-vax study, a big newspaper writing an article against it etc.), depending on the amount of infected individuals I. Taking all of these thoughts into consideration, we propose the following transition rates:

Model 4.1 (Stochastic Process within the Combined Model). Consider the same context as in Model 3.4. Additionally, let $n_X, n_Y, a \in \mathbb{N}$ be scaling factors and

 $\Omega(I): [0,1] \to \mathbb{R}^+$. They all describe different influences on the vaccination behaviour of the population and let I be the number of infected individuals as defined in (1). Then, the rates of the underlying stochastic process read as:

$$X_t \to X_t + 1 \text{ at rate } \mu \left(N - X_t \right) \frac{\theta_X (X_t + N_X + aI)}{\theta_X (X_t + N_X + aI) + \left(N - X_t + N_Y + \Omega \left(I \right) N \right)}$$
(10)
$$\theta_Y \left(N - X_t + N_Y + \Omega \left(I \right) N \right)$$

$$X_t \to X_t - 1 \text{ at rate } \mu X_t \frac{\theta_Y \left(N - X_t + N_Y + \Omega\left(I\right) N \right)}{\left(X_t + N_X + aI \right) + \theta_Y \left(N - X_t + N_Y + \Omega\left(I\right) N \right)}$$

By following common *SIR*-modeling approaches with standard incidence, we can write down a first draft of our model:

$$\dot{S} = -\beta S \frac{I}{N} + bN \left(1 - \frac{X_t}{N} \right) - \delta S$$

$$\dot{I} = \beta S \frac{I}{N} - \alpha I - \delta I$$

$$\dot{R} = \alpha I + bN \frac{X_t}{N} - \delta R$$

(11)

This only takes into account the susceptible, infected and recovered individuals, but we also want to consider the amount of vaccinated subjects, which is currently described by a stochastic process:

 $X_t \to X_t \pm 1$ with rates as described in (10)

In this context, β is the rate of infection and α the recovery rate, b stands for the birth and δ for the death rate of the population. As one can see, this is a stochastic process mixed up with a three-dimensional system of ODEs, leading to a fourth dimension as soon as X_t will be included. But we want to work solely with a system of ODEs, that has less than four dimensions preferably. To do so, we will have to compute the deterministic limit of all of the components of the model $(S, I, R \text{ and } X_t)$. Additionally, we will use $\dot{N} = bN - dN$ (another SIR-model commonalty). We define:

$$s \coloneqq \frac{S}{N}, i \coloneqq \frac{I}{N}, r \coloneqq \frac{R}{N}, x \coloneqq \frac{X_t}{N}$$
(12)

We may now compute:

$$\dot{s} = \frac{N\dot{S} - S\dot{N}}{N^2} = \frac{\dot{S}}{N} - \frac{S}{N}\frac{\dot{N}}{N}$$
$$\stackrel{(11,12)}{=} -\beta si + b\big((1-x) - s\big)$$
$$\dot{i} = \frac{N\dot{I} - I\dot{N}}{N^2} = \frac{\dot{I}}{N} - \frac{I}{N}\frac{\dot{N}}{N}$$
$$\stackrel{(11,12)}{=} \beta si - i(\alpha + b)$$
$$\dot{r} = \frac{N\dot{R} - R\dot{N}}{N^2} = \frac{\dot{R}}{N} - \frac{R}{N}\frac{\dot{N}}{N}$$

$$\stackrel{(11,12)}{=} \alpha i + bx - rb$$

Note that \dot{r} is not needed in the model any more, since it occurs in none of the other equations (it will be shown in Subsection 4.2 that it does not occur in \dot{x} neither), leading to our model being four-dimensional, but virtually three-dimensional! This will facilitate the upcoming model analysis tremendously.

4.2 Computation of the Deterministic Limit

In addition to the context of Model 4.1, we will assume that the two zealot groups N_X and N_Y scale linearly with the total population size N:

Proposition 4.1. Let $i \coloneqq \frac{I}{N}$, $n_X \coloneqq \frac{N_X}{N}$, $n_Y \coloneqq \frac{N_Y}{N}$ and $\omega(i) \coloneqq \frac{\Omega(I)}{N}$. Then, the deterministic limit for $x(t) \coloneqq \frac{X_t}{N}$ reads

$$\dot{x} = -\mu x \frac{\theta_Y \left(1 - x + n_Y + \omega\left(i\right)\right)}{\left(x + n_X + ai\right) + \theta_Y \left(1 - x + n_Y + \omega\left(i\right)\right)} + \mu \left(1 - x\right) \frac{\theta_X \left(x + n_X + ai\right)}{\theta_X \left(x + n_X + ai\right) + \left(1 - x + n_Y + \omega\left(i\right)\right)}$$
(13)

The proof of the claim above will be quite lengthy, but mainly computational. Firstly, we will state the master equations of our model, i.e. a system of first-order ODEs which describes how our modelled system behaves over time in a probabilistic way. In other words: the system is in state i with probability p_i and the master equations are differential

equations of those p_i . Afterwards, we will make use of these and compute the Kramers-Moyal expansion: a generalisation of the Fokker-Planck equation that arises. Both of these concepts are well explained in [49]:

Definition 4.1 (Fokker-Planck Equation for One Variable by [49]). Let B(v,t) be the distribution function for one-dimensional Brownian motion and $D^{(1)}(x)$, $D^{(2)}(x)$ with $D^{(2)}(x) > 0$ the drift and diffusion coefficient, respectively. Both of them may also depend on time. Then, the general Fokker-Planck equation for one variable x reads:

$$\partial_t \left(B\left(v,t\right) \right) = \left[-\partial_x \left(D^{(1)}(x) \right) + \partial_x^2 \left(D^{(2)}(x) \right) \right] B\left(v,t\right).$$

Definition 4.2 (Kramers-Moyal Expansion by [49]). The general expansion of a Fokker-Planck equation that does not stop after the second derivative and also contains higher ones with respect to x, has infinite number of terms and is called the Kramers-Moyal expansion. It reads:

$$\partial_t \left(B\left(v,t\right) \right) = \sum_{i=1}^{\infty} \left(-\partial_x \right)^i \left(D^{(i)}(x) B\left(v,t\right) \right).$$

One will notice whilst reading the proof below that the Kramers-Moyal expansion which will be computed is only a sum of two, not infinitely many components. This is due to rigorous arguments such as Langevin equations and Gaussian δ -correlated noise, which would blow up the scope of this work. To gain any more insights on those arguments and the theory behind Fokker-Planck Equations, how they arise and what their purpose is, we highly encourage reading Risken's book [49]. We may now state the proof of Proposition 4.1:

Proof. For better readability and improved simplicity, we will refer to the transition rates of the model as:

$$F_{+}(X_{t}) \coloneqq \frac{\theta_{X}(X_{t} + N_{X} + aI)}{\theta_{X}(X_{t} + N_{X} + aI) + (N - X_{t} + N_{Y} + \Omega(I)N)}$$

$$F_{-}(X_{t}) \coloneqq \frac{\theta_{Y} \left(N - X_{t} + N_{Y} + \Omega\left(I\right) N \right)}{\left(X_{t} + N_{X} + aI \right) + \theta_{Y} \left(N - X_{t} + N_{Y} + \Omega\left(I\right) N \right)}$$

Having clarified the notation, the master equation now reads:

$$\dot{p}_{k}(t) = F_{+}(k-1) \times p_{k-1}(t) + F_{-}(k+1) \times p_{k+1}(t) - \left(F_{+}(k) + F_{-}(k)\right) \times p_{k}(t)$$
(14)

In the next step, we will compute the Kramers-Moyal expansion of (14). To do so, we assume:

$$p_k(t) \approx hu\left(\frac{k}{N}, t\right), h = \frac{1}{N} \text{ and } x = hk.$$
 (15)

Again, we define for simplicity by expanding $F_+(x)$ and $F_-(x)$ with $\frac{h}{h}$, respectively:

$$f_{+}(x) \coloneqq \frac{\theta_{X}(x+n_{X}+ai)}{\theta_{X}(x+n_{X}+ai) + (1-x+n_{Y}+\omega(i))}$$

$$f_{-}(x) \coloneqq \frac{\theta_{Y}(1-x+n_{Y}+\omega(i))}{(x+n_{X}+ai) + \theta_{Y}(1-x+n_{Y}+\omega(i))}$$
(16)

We compute:

$$\partial_{t} \left(u \left(\frac{k}{n}, t \right) \right) = \frac{1}{h} \dot{p}_{k} \left(t \right)$$

$$\stackrel{(14,15)}{=} \frac{1}{h} \left[F_{+} (k-1) h u (x-h,t) + F_{-} (k+1) h u (x+h,t) - \left(F_{+} (k) + F_{-} (k) \right) h u (x,t) \right]$$

$$= \frac{1}{h} \left[\mu \left(1 - (x-h) \right) f_{+} (x-h) u (x-h,t) + \mu (x+h) f_{-} (x+h) u (x+h,t) - \left(\mu (1-x) f_{+} (x) + \mu x f_{-} (x) \right) u (x,t) \right]$$

$$(17)$$

From there we can compute the Taylor series of second order for the x - h and x + h terms:

$$\mu (1 - (x - h)) f_{+}(x - h) u (x - h, t)$$

= $\mu (1 - x) f_{+}(x) u (x, t) - h \partial_{x} (\mu (1 - x) f_{+}(x) u)$
+ $\frac{1}{2} h^{2} \partial_{x}^{2} (\mu (1 - x) f_{+}(x) u)$ (18)

$$\mu(x+h)f_{-}(x+h)u(x+h,t) = \mu x f_{-}(x)u(x,t) + h\partial_{x}(\mu x f_{-}(x)u) + \frac{1}{2}h^{2}\partial_{x}^{2}(\mu x f_{-}(x)u)$$

We may now resume the computations (17) from above, add those new results and compute the Fokker-Planck equation:

$$\partial_t \left(u\left(\frac{k}{n}, t\right) \right) = \dots$$

$$\stackrel{(15,18)}{=} -\partial_x \left(\mu u(x,t) \left[(1-x)f_+(x) - xf_-(x) \right] \right) + \frac{1}{2N} \partial_x^2 \left(\mu u(x,t) \left[(1-x)f_+(x) + xf_-(x) \right] \right)$$
(19)

So, for $N \to \infty$:

$$\partial_t \left(u\left(\frac{k}{n}, t\right) \right) = -\partial_x \left(\mu u(x, t) \left[(1-x)f_+(x) - xf_-(x) \right] \right)$$
(20)

By the definition in (16), this means that the ODE due to the drift term in case $N \to \infty$ reads

$$\dot{x} = -\mu x \frac{\theta_Y \left(1 - x + n_Y + \omega \left(i\right)\right)}{\left(x + n_X + ai\right) + \theta_Y \left(1 - x + n_Y + \omega \left(i\right)\right)} + \mu \left(1 - x\right) \frac{\theta_X \left(x + n_X + ai\right)}{\theta_X \left(x + n_X + ai\right) + \left(1 - x + n_Y + \omega \left(i\right)\right)}$$

Which is exactly the result stated in (13)

To get a better understanding of the dynamics that are hidden in those equations, we propose to take a look at Figure 2. The interactions should then be clear and especially the fact, that there is no interaction between neither s and r, nor i and r, which is the reason r does not appear in the definition of Model 4.2.

Figure 2: Dynamics of the Combined Model



4.3 Definition of the Combined Model

Now that all the computations are done, we may regroup all of the results and formulate the system of ODEs that describes our model:

Model 4.2 (Combined Model). Consider a population of size $N \in \mathbb{N}$, where $X_t \in \{0, \ldots, N\}$ and $N - X_t$ describe the amount of supporters of opinion X and Y at time t, respectively. Furthermore, let N_X and $N_Y > 0$ be the number of zealots that support X and Y, respectively. Let $\alpha, \beta, a, b \in \mathbb{N}$, with the same meaning as in (11) and ω (i) just as in (13). An individual of the population (who is not a zealot) may reconsider his opinion with rate $\mu \in \mathbb{N}$ via the same process as described in Model 3.3 and in addition of the contact rates $\theta_X, \theta_Y \in \mathbb{N}$. Let $S \in \{0, \ldots, N\}$ and $I \in \{0, \ldots, N\}$ describe the amount of susceptible and infected individuals, just as in (1). Now, rescale S, I and X_t as done in (12) and assume N_X and N_Y to scale linearly with the population size N, as in Proposition 4.1. Then, the model reads as follows:

$$\dot{s} = -\beta si + b((1-x) - s)$$

$$\dot{i} = \beta si - i(\alpha + b)$$

$$\dot{x} = -\mu x \frac{\theta_Y (1 - x + n_Y + \omega(i))}{(x + n_X + ai) + \theta_Y (1 - x + n_Y + \omega(i))}$$

$$+ \mu (1-x) \frac{\theta_X (x + n_X + ai)}{\theta_X (x + n_X + ai) + (1 - x + n_Y + \omega(i))}$$
(21)

4.4 Bifurcation Analysis

This Subsection is dedicated to the analysis of the proposed model in order to understand it better and to get a glimpse at what it is capable of. A major part of this is the determination of stationary points, also known as equilibrium points, and bifurcation analysis where we will try to find all the bifurcations the system undergoes. As far as equilibria go, they should be a familiar concept from the theory of ODEs. Consider $f(x) \in C^1(\mathbb{R}^n, \mathbb{R}^n)$ and $\dot{x} = f(x)$. A point \hat{x} with $f(\hat{x}) = 0$ is called an equilibrium point. Regarding bifurcation analysis, this work will not be able to give a full introduction to it since it represents an extensive topic. For everybody interested in diving deeper into bifurcations and model dynamics, we recommend reading Kuznetsov's, as well as Hale and Koçak's books [50, 51]. We will restrict ourselves on briefly presenting the bifurcations mentioned in this work:

Definition 4.3 (Pitchfork Bifurcation in 1D by [21]). Consider $x(t) : \mathbb{R} \to \mathbb{R}$. Let a stationary point exist that is stable for parameter values $\mu \leq \mu_0$. Additionally, let two new stationary points show up for parameter values $\mu > \mu_0$ and let those two new points be stable, whilst the original stationary point becomes unstable. So, there is a Pitchforkbifurcation at the parameter value μ_0 and its normal form reads:

$$\dot{x} = \mu x - x^3 \tag{22}$$

Definition 4.4 (Saddle-Node Bifurcation in 1D by [21]). Consider $x(t) : \mathbb{R} \to \mathbb{R}$. A bifurcation with bifurcation parameter μ and normal form

$$\dot{x} = \mu - x^2 \tag{23}$$

is called a saddle-node bifurcation.

Note that the normal form of the saddle-node bifurcation implies no existing fixed points for $\mu < 0$, exactly one for $\mu = 0$ and two for $\mu > 0$. This is why they are also called "blue-sky-bifurcations", as they seem to appear "out of nowhere" [21].

Definition 4.5 (Hopf Bifurcation by [51]). Let $z : \mathbb{R}^n \to \mathbb{R}^n$ with $n \ge 2$. Consider the continuous dynamic system $\dot{z} = f(z)$ and its linearisation, described by its Jacobian J(z). Let \hat{z} be a stationary point of the system and consider the linearisation $J_{\hat{z}} := J(\hat{z})$ at that point. If

$$\forall \lambda \in \operatorname{spec} \left(J_{\hat{z}} \right) \setminus \{ \lambda^*, \overline{\lambda^*} \} : Re(\lambda) < 0 \land Re(\lambda^*) = 0$$

holds, then a Hopf-bifurcation occurs as soon as the pair $\lambda^*, \overline{\lambda^*}$ crosses the imaginary axis due to a bifurcation parameter change.

We may now begin with the analysis of our system.

After trying several different approaches, we settled on the assumption to fix the "antivax influence": $\omega(i) := a(1-i)$. With this adjustment, the model indicates a tendency towards vaccination for increasing i and a tendency towards not getting vaccinated for decreasing i, which we esteem to be realistic. By additionally uncoupling \dot{x} and \dot{i} , such that the number of infected individuals has no more influence on the number of provaxxers, we even know about two bifurcations in the model now! This can be achieved by setting a = 0 and falling back to the model presented by Müller and Tellier in [47]. As shown there, the following proposition may be formulated:

Proposition 4.2 (Pitchfork Bifurcation for Special Case by [47]). Consider Model 4.2. For $n_X = n_Y = n$, $\theta_X = \theta_Y = \theta_P$ and a = 0, $x = \frac{1}{2}$ is always a stationary point and shows a pitchfork bifurcation at θ_P with

$$\theta_P = \frac{1-2n}{1+2n}$$

Before we begin with the proof, we would like to remark that s, as well as i have not been mentioned in Proposition 4.2, since Müller and Tellier analysed a one-dimensional system in their work. But as will be seen below, we can easily construct a stationary point (s^*, i^*, x^*) by setting s^* and i^* as in (24) and (25) and $x^* = \frac{1}{2}$.

Proof. This can quickly be shown by computing the Taylor expansion of third degree at the point $x = \frac{1}{2}$ of (13)'s right-hand-side. When replacing θ_X and θ_Y with θ_P , one will see that the linear term disappears, leaving only the term of third order and the rest term. Thus, following the normal form (22), the presence of a pitchfork bifurcation at that point has been shown. For the detailed proof see [47].

As also shown by Müller and Tellier, the pitchfork bifurcation changes into a saddle-node bifurcation for slightly differing numbers of zealots n_X and n_Y .

Writing down a general stationary point analytically for our model turns out to be very unpractical, but we are able to produce a similar result to the one presented in Proposition 4.2, without having to restrict a = 0:

Proposition 4.3 (A Stationary Point of the Combined Model). Presume (s^*, i^*, x^*) , $i^* \neq 0$ is a stationary point of Model 4.2. For $n_X = \max\{a (1 - 2i^*), 0\}$, $n_Y = a (2i^* - 1) + n_X$ and $\theta_X = \theta_Y = \frac{1-2n}{1+2n}$ with $n \coloneqq ai^* + n_X$, there always is a stationary point at $x^* = \frac{1}{2}$.

Proof. We set $x^* = \frac{1}{2}$. It is now very easy to write down s^* and i^* :

$$0 \stackrel{!}{=} s^{*} \stackrel{(4.2)}{=} -\beta s^{*} i^{*} + b \left(\frac{1}{2} - s^{*}\right)$$

$$\Leftrightarrow s^{*} = \frac{b}{2\beta i^{*} + 2b}$$
(24)

$$0 \stackrel{!}{=} i^{*} \stackrel{(4.2)}{=} \beta s^{*} i^{*} - \alpha i^{*} - b i^{*}$$

$$\stackrel{i^{*} \neq 0}{\Leftrightarrow} i^{*} = \frac{b}{2(\alpha + b)} - \frac{b}{\beta} \stackrel{(24)}{\Rightarrow} s^{*} = \frac{\alpha + b}{\beta}$$
(25)

So we have a stationary point at $\left(\frac{\alpha+b}{\beta}, \frac{b}{2(\alpha+b)} - \frac{b}{\beta}, \frac{1}{2}\right)$. Now we have to adjust the parameters occurring in (13) such that $\dot{x} = 0$ holds. To do so, we will compare the two parts of \dot{x} and make sure they are the same, which will make them cancel each other out and yield the wanted result. We immediately see that $\mu x = \mu (1-x)$ for $x = x^* = \frac{1}{2}$. We may now write down two equations that have to hold to ensure a stationary state at x^* :

$$\theta_X\left(\frac{1}{2} + n_x + ai^*\right) = \theta_Y\left(\frac{1}{2} + n_Y + a(1 - i^*)\right)$$
(26)

$$\theta_X \left(\frac{1}{2} + n_X + ai^*\right) + \frac{1}{2} + n_Y + a(1 - i^*) = \frac{1}{2} + n_X + ai^* + \theta_Y \left(\frac{1}{2} + n_Y + a(1 - i^*)\right)$$
(27)

One quickly notices that both these equations hold true for $n_X = n_Y = n$, $\theta_X = \theta_Y = \theta$ and $i^* = \frac{1}{2}$, but we do not want to fix i^* . So we have to look at another constellation of parameters, that ensures (26) and (27) both hold. We find the following two possibilities:

Constellation 1	Constellation 2
$n_X = 0$	$n_X = a(1 - 2i^*)$
$n_Y = a(2i^* - 1)$	$n_Y = 0$
$\theta_X = \theta_Y = \theta$	$\theta_X = \theta_Y = \theta$
$\theta = \frac{1 - 2ai^*}{1 + 2ai^*}$	$\theta = \frac{1 + 2ai^* - 2a}{1 + 2a}$

As n_X of Constellation 2 and n_Y of Constellation 1 could be less than zero for some parameters, we formulate a general definition for n_X and n_Y , which is equivalent to the results above:

$$n_X = \max\{a(1-2i^*), 0\} n_Y = a(2i^*-1) + n_X$$
(28)

Furthermore, we can rewrite both θ into one general formula as well:

$$\theta \coloneqq \frac{1-2n}{1+2n} \text{ with } n \coloneqq ai^* + n_X \tag{29}$$

Thus, $\dot{x} = 0 \Leftrightarrow (28) \land (29)$, which is exactly what Proposition 4.3 claims.

Furthermore, there are multiple indicators that imply the existence of a Hopf bifurcation in the system. For further investigations in that directions, we will use the insights about the theoretical stationary point $(s^*, i^*, 0.5)$ from Proposition 4.3 and set the parameters of our model as specified in Table 1.

Parameter	Value
α	0.45
eta	5
μ	23.05
a	0.1
b	2
n_X	$\max\{a\left(1-2i^*\right),0\}$
n_Y	$a\left(2i^*-1\right)+n_X$
θ_X	$\frac{1-2(ai^*+n_X)}{1+2(ai^*+n_X)}$
$ heta_Y$	$ heta_X$

Table 1: Model context for strong negative correlation

Now, define the function $antivax(x, i) : [0, 1]^2 \to \mathbb{R}$ so that it describes the anti-vaccination impact on \dot{x} and plot it against x. To do so, we define antivax(x, i) to be the negative term of the right hand side in (13), when leaving out μ :

$$antivax(x,i) \coloneqq -x \frac{\theta_Y (1 - x + n_Y + \omega(i))}{(x + n_X + ai) + \theta_Y (1 - x + n_Y + \omega(i))}$$

As one can see quite well in Figure 3, very little changes to i have a massive impact on x. The black curve depicts the anti-vax impact when plugging in the i-value of the theoretical stationary point we described in Proposition 4.3. The two upper blue curves show



Figure 3: Anti-Vax Impact vs. Number of Pro-Vaxxers

the effect of plugging in $i^* + 0.01$ and $i^* + 0.1$, while the two lower red curves were given $i^* - 0.01$ and $i^* - 0.1$, respectively. As one can see, a small change $i^* \pm \epsilon$ with $\epsilon = 0.01$ already made the *x*-component slide to 0.4 and 0.6, respectively, whereas $\epsilon = 0.1$ caused a way more considerable translation towards 0.76 and 0.24, respectively. This leads us to the conclusion that there is a strong negative correlation between *x* and *i*, which is necessarily needed for a Hopf bifurcation.

Furthermore, when we take a look at the system in all its entirety and let the simulation begin at $s_0 = 0.8$, $i_0 = 0.2$, $x_0 = 0.2$, oscillations of s, x and i appear. Note that i has been scaled to 10i in Figure 4, so that the oscillations of i become visible.

As we know from Definition 4.5, a Hopf bifurcation occurs when the two purely imaginary eigenvalues cross the imaginary axis. This fact can be seen in Figure 5, which depicts the spectrum of the linearisation J_{μ} of our model, with respect to μ . Note that the spectrum seems to only consist of multiple complex conjugated eigenvalue pairs $\lambda_{1,\mu}$ and $\lambda_{2,\mu} = \overline{\lambda_{1,\mu}}$. This is due to the fact that the third eigenvalue $\lambda_{3,\mu}$ (as our model is three-dimensional) has always been purely real, with $\lambda_{3,\mu} < 1$, leading to the complex eigenvalues being indistinguishable when looking at the entire spectrum. As we are only interested in the complex ones, we zoomed into the spectrum to render the complex eigenvalues more visible, which is why none of the $\lambda_{3,\mu}$ can be seen in Figure 5.



Figure 4: Behaviour Over Time of the Combined Model





We computed the deterministic limit of Model 4.2 in Subsection 4.2 so that we could analyse its behaviour over time and visualise it quite nicely. We now intend to fit the model onto real-life data and to see how well it performs. It has been shown in [52, 53], that weak effect models are better suited for this goal, especially considering realism, which is why the next step will be to compute the weak effect limit of our model.

Theorem 4.1 (Weak Effect Limit of the Combined Model). Let $N_i, i \in \{X, Y\}$ denote the numbers of zealots of opinion X and Y, respectively and N the population size. Furthermore, let the reinforcement within the model be described by $\theta_i = 1 - \frac{\Theta_i}{N}$. For $N \to \infty$, the density of the invariant measure for the random variable $z_t = \frac{X_t}{N}$ is given by

$$\phi(x) = C e^{\frac{1}{2}(\Theta_X + \Theta_Y)x^2 - \Theta_X x} x^{N_X + Ai - 1} \left(1 - x\right)^{N_Y + A - Ai - 1}$$
(30)

where C is determined by the prerequisite $\int_0^1 \phi(x) dx = 1$.

Proof. We take a look back at (19), the Fokker-Planck equation for which we computed the Kramers-Moyal expansion before. There, the scaling was $n_i = N_i/N$ and θ_i constant in N. Afterwards, we will use a slightly different scaling which we will prefer for the Dirichlet limit. We recall the transition rates from (16):

$$f_{+}(x) \coloneqq \frac{\theta_{X}(x + n_{X} + ai)}{\theta_{X}(x + n_{X} + ai) + (1 - x + n_{Y} + a(1 - i))}$$

$$f_{-}(x) \coloneqq \frac{\theta_{Y}(1 - x + n_{Y} + a(1 - i))}{(x + n_{X} + ai) + \theta_{Y}(1 - x + n_{Y} + a(1 - i))}$$
(31)

Thus, the limiting Fokker-Planck equation reads:

$$\partial_t (u(x,t)) = -\partial_x (\mu u(x,t) ((1-x) f_+(x) - x f_-(x))) + \frac{1}{2N} \partial_x^2 (\mu u(x,t) ((1-x) f_+(x) + x f_-(x)))$$
(32)

By rewriting drift and noise term with $n_i = N_i/N$, $\theta_i = 1 - \frac{\Theta_i}{N}$, h = 1/N, a = A/N and neglecting terms of $\mathcal{O}(N^{-2})$ we find (using maxima [54]):

$$\mu(1-x)f_{+}(x) - \mu x f_{-}(x)$$

$$= \mu(1-x)\frac{(1-h\Theta_{X})(x+hN_{X}+hAi)}{(1-h\Theta_{X})(x+hN_{X}+hAi) + (1-x+hN_{Y}+hA(1-i))}$$

$$- \mu x \frac{(1-h\Theta_{Y})(1-x+hN_{Y}+hA(1-i))}{(x+h_{N}X+hAi) + (1-h\Theta_{Y})(1-x+hN_{Y}+hA(1-i))}$$

$$= \mu \Big(\left[(\Theta_{X}+\Theta_{Y})x - \Theta_{X} \right] x (1-x) - (N_{X}+N_{Y}+A)x + Ai + N_{X} \Big) h + \mathcal{O}(h^{2}),$$

while

$$h[\mu(1-x)f_{+}(x) + \mu x f_{-}(x)] = \mu h 2x(1-x) + \mathcal{O}(h^{2}).$$

After rescaling time: $T = \mu ht$, the Fokker-Planck equation (32) becomes

$$\partial_T \left(u \left(x, T \right) \right)$$

= $-\partial_x \left(u(x,T) \left(\left[\left(\Theta_X + \Theta_Y \right) x - \Theta_X \right] x \left(1 - x \right) - \left(N_X + N_Y + A \right) x + Ai + N_X \right) \right)$
+ $\partial_x^2 \left(u(x,T) x (1 - x) \right)$

For the invariant distribution $\phi(x)$, the flux of the rescaled Fokker-Planck equation is zero, that is:

$$0 = -\left(\left[\left(\Theta_X + \Theta_Y\right)x - \Theta_X\right]x\left(1 - x\right) - \left(N_X + N_Y + A\right)x + Ai + N_X\right)\phi(x) + \frac{d}{dx}\left(\phi(x)x\left(1 - x\right)\right)$$

With $v(x) = \phi(x)x(1-x)$, we have:

$$v'(x) = \left(\left[\left(\Theta_X + \Theta_Y \right) x - \Theta_X \right] + \frac{N_X}{x} - \frac{N_Y}{1 - x} + \frac{A(i - x)}{x(1 - x)} \right) v(x)$$

and hence

$$v(x) = Ce^{\frac{1}{2}(\Theta_X + \Theta_Y)x^2 - \Theta_X x} x^{N_X + Ai} (1 - x)^{N_Y + A - Ai}$$

Respectively,

$$\phi(x) = C e^{\frac{1}{2}(\Theta_X + \Theta_Y)x^2 - \Theta_X x} x^{N_X + Ai - 1} (1 - x)^{N_Y + A - Ai - 1}$$

With the weak effect limit in hand, we would now like to analyse real-life data. To do so, we retrieved the vaccination coverage against Meningococcus C in German counties between 2009 and 2014 from [55] and the matching incidences per county of the disease from [56]. The next step is to generate values for all of the parameters that occur in (30), via maximum-likelihood-estimation. The integral that defines C becomes numerically unstable for large parameter values, so we re-parametrise the distribution, defining $\hat{\nu}, \hat{s}, \hat{\theta}, \hat{\psi}, \hat{\xi}$ by

$$\Theta_X = \hat{s}\hat{\theta}\hat{\psi}$$

$$\Theta_Y = \hat{s}\hat{\theta}(1-\hat{\psi})$$

$$N_X - 1 = \hat{s}(1-\hat{\theta})\hat{\nu}(1-\hat{\xi})$$

$$N_Y - 1 = \hat{s}(1-\hat{\theta})(1-\hat{\nu})(1-\hat{\xi})$$

$$A = \hat{s}\hat{\theta}\hat{\xi}$$

where $\hat{s}, \hat{\theta}, \hat{\nu}, \hat{\xi} \in [0, 1]$ and $\hat{s} > 0$, with the restriction $\hat{s}(1 - \hat{\theta})\hat{\nu}(1 - \hat{\xi}) > 1$ and $\hat{s}(1 - \hat{\theta})(1 - \hat{\nu})(1 - \hat{\xi}) > 1$. Therewith, the distribution becomes

$$\phi(x) = \hat{C} \exp\left[\hat{s}\left\{\hat{\theta}\left(\frac{x^2}{2} - \hat{\psi}x\right) + \ln(x)\left((1 - \hat{\theta})\hat{\nu}(1 - \hat{\xi}) + \hat{\theta}\hat{\xi}i\right) + \ln(1 - x)\left((1 - \hat{\theta})(1 - \hat{\nu})(1 - \hat{\xi}) + \hat{\theta}\hat{\xi}(1 - i)\right) + B\right\}\right],$$
(33)

where B is a constant depending on the data at hand and chosen in such a manner, that it avoids an exponent with very large absolute number. The constant \hat{C} is, as was Cbefore, determined by the integral being equal to 1.

We then proceeded with "cleaning" the data, which in our case meant to remove all vaccination coverage data points with a value below 0.5 and renormalise the data points which were only in the interval $\left[\frac{1}{2},1\right]$ afterwards, so that we stay in the interval $\left[0,1\right]$ and keep the initial ordering. This step is justified by our assumption that 50% of the individuals are not ready to change their opinion and a discussion about taking a shot or not only emerges in the other 50% of the population. Additionally, many German counties did not have registered incidence values, so we replaced all "non-available" data points with a dummy value. We set the dummy to be the mean of all available incidence values, which turned out to be approximately 0.0225. During the first approaches, we used every county's own incidence value i_{county} to compute every county's distribution function, leading to every iteration step producing n = 399 curves. We found that all curves were very close to each other, thus the minimal differences between all the incidence values that happened in the low one-digit range did not affect the optimisation result very much. So for simplicity's sake we decided to recompute the new mean of all incidences (after all "non-available" data points had been filled up with the dummy) and to use this as a general incidence value for all counties. After that, we ran our maximum-likelihood estimation R-Script and performed a Kolmogorov-Smirnov test on the optimised model with initial parameters $\hat{s}_{init} = 100, \hat{\nu}_{init} = \hat{\theta}_{init} = \hat{\psi}_{init} = \hat{\xi}_{init} = 0.5$. One may have noticed that vaccination data, especially the one we used, is discrete and our model, as well as the Kolmogorov-Smirnov test, expect continuous values. This led to the occurrence of ties in the data set. The optimisation consisted of three runs. First, we set both of the reinforcement parameters to zero and fell back to the zealot model. Then, we set both reinforcement parameters to be equal and optimised the equal reinforcement model. Finally, we let the reinforcements parameters vary freely. The results of our optimisation can be seen in Figures 6a-c.

We notice that the zealot and equal reinforcement model fit the data in a very similar way, but both not very well. The Kolmogorov-Smirnov tests for both of these models yielded p-values in $\mathcal{O}(10^{-6})$, thus they could not be accepted as models of the data. On the other hand, the full reinforcement model fits remarkably well. The Kolmogorov-Smirnov test also yielded a p-value of 0.73 for the full-reinforcement model, so we may accept this model and our distribution (33) as the distribution of the vaccination data.

Of course we do not claim our model to be exhaustive and as is well-known: "no model is right". All of the comments, assumptions and claims made in the following base on the results our work yielded. We did not take a look at additional cofactors, such as educational or social background or financial status. These are all possible determinants of behaviour, just as communication and filter bubbles and represent interesting research topics for further work.



Figure 6: Data Fits of the Three Different Models

5 Discussion

Despite the vast amount of studies and publications that show how diseases can be eradicated by vaccination, the large number of casualties that occur in communities with extremely low vaccination coverage or the practical fact that diseases such as Poliomyelitis in Europe do not represent any dangers to life any more, vaccination hesitancy is still undermining mankind's fight against diseases. Classified as one of the Top Ten Threats to global health in 2019, vaccine hesitancy is the main obstacle between our current situation and a world with less disease [41]. Even in Germany, where the percentage of anti-vaxxers is quite low (3-5% according to [22]) herd immunity is by far not reached for every disease there is a vaccine for. Figure 7 illustrates the example of Meningococcus C and its immunisation rates. As per [57], the required vaccination coverage needed for herd immunity depends on the infectivity of the disease and its reproduction number R_0 , but is generally reached at about 75 - 95% vaccination coverage. The actual coverage in Germany (from 2009 to 2014) is outside of that range for approximatively half of all districts, especially in the south-east, where some regions show an absolute low record of 31,62 - 44,24%.



Figure 7: Immunisation Against Meningococcus C in Germany - Implementation Analysis in 2013 of the STIKO Recommendations from 2009 to 2014 [55]

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To understand the dynamics of this paradoxical behaviour, this work has taken into consideration the work done by Bauch and Bhattacharyya in [20] as well as by Müller and Tellier in [47]. The aim was to join the behaviour-incidence approach and the zealot model with reinforcement to render the compartmental model of Bauch more realistic when depicting the behaviour of individuals regarding vaccination by adding communication between individuals and an opinion formation process. Similar work has been done by Tyson et al. in [58], where the mentioned SIR-Model and its disease dynamics has been combined with opinion dynamics. Tyson et al. stressed the fact that their model, does not work with utility or contagion in mind, but emphases the impact of communication and persuasion. In our model, the aspect of opinion formation and influence through fellow individuals, social media and filter bubbles, following the ansatz of pay-off functions, was added to the underlying stochastic process before computing the model dynamics with the deterministic limit. By doing so, we found that the dynamics are completely disjoint from the number of recovered individuals. Furthermore, we found several bifurcations which appear in the model. For a = 0 the model falls back to the results presented in [47] and undergoes a pitchfork bifurcation, as well as a saddle-node bifurcation. The presence of a Hopf bifurcation in our system has been indicated by Figures 4 and 5. It can be stated explicitly with the help of different tools, e.g. xppaut [59]. This insight shows evidence of oscillations in the system, as can be seen in Figure 4. It would be very interesting for future work to investigate the existence of a Takens-Bogdanov bifurcation in the presented system.

We found a stationary point for a fixed $x = \frac{1}{2}$ in our system, as well as a strong correlation between x and i shown in Figure 3. The oscillations of x and i express the varying immunisation behaviour within the population: when the amount of infected individuals is low, the disease is not strongly perceived by the population, thus the immunisation rate drops. The small number of immune subjects will then over time generate a high rate of infected subjects, which leads to increased disease awareness within the population and higher immunisation rates. The cycle we can observe here can be interpreted in a certain manner as the biological equivalent to the "pork cycle" or "cattle cycle" known from economics [60].

Dong et al. complained about the fact that most studies concerning opinion formation and opinion dynamics focused on the analysis of simulations, generated with random data [61]. So in Subsection 4.5, we aimed at fitting our model onto the real-life data retrieved from [55], visualised in Figure 7 above, and from [56]. As mentioned there, taking into account every single incidence value of every single county did not have a noticeable impact on the optimisation results compared to using the mean as a general incidence value. This shows that little changes to the incidence rate do not provoke any changes to the vaccination rate. Speaking more generally: a small amount of new infected individuals or a small decrease of the disease's prevalence will not lead to a change of vaccination behaviour in the population. This leads us to the assumption that the increase or decrease of infections has to surpass a certain threshold before impacting behaviour, thus the communication between the population has a bigger impact on behaviour, than the rate of incidence. Furthermore, the very small p-value for the equal reinforcement model indicates a larger amount of reinforcement usage for the anti-vaxxers, than for the provaxxers. This matches our expectations and the analogous findings from [47] regarding populist parties, which are considered to be more firm in their views compared to other parties.

Filter bubbles and communication seem to play a major role in the dynamics of vaccination behaviour. An interesting next step would be to reflect on how to minimise or even destroy these bubbles. As our data showed, the bubbles appear locally, in small subsets of the entire population, i.e. in counties. A possible measure could be to invest more in local, small-scaled politics, visit the counties to propagate the scientific findings and knowledge about the vaccination process and do educational work considering vaccines. A considerable amount of work has still to be done in this area, in order to raise awareness, reduce the spread of fake-news and eventually, ban vaccine hesitancy from the top ten threats to global health.

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6 Appendix

6.1 R-Script for Figures 3, 4 and 5

```
#
    \# SIRS model with reinforcement
    #
3
    # Aim:
    # find a Hopf bifurcation
    #
    # parameter
7
                   \# contact rate
    bbeta = 5
    alpha = 0.45 \# recovery rate
9
          = 2
                   \# birthrate – determines the time scale — fixed.
    B
    # we calibrate the system s.t. x=1/2 is a stat states
    s.star = (B+alpha)/bbeta;
13
    i.star = (1-0.5) *B/(alpha+B)-B/bbeta;
    \operatorname{cat}("i.star = ", i.star, "\backslash n");
15
    theta1 = 1; \# reinforcement parameter of pro-vax (x)
17
    theta2 = theta1; \# reinforcement parameter of anti-vax (1-x)
            = 10;
                      # zealots pro vax
    n1
19
    n2
            = 10:
                      # zealots anti-vax
            = 0.1;
                       # influence I
21
    \mathbf{a}
            = 200 \# 1;
                         # time scale reinforce
    с
23
    \# adapt n1, n2 s.t. a*i+n1 = a*(1-i)+n2
    \# choose n1 minimal, s.t. we have a non-negative n2 (=0)
25
    n1 = max(c(a*(1-2*i.star), 0));
    n2 = a * (2 * i . star - 1) + n1;
27
    n.star = a*i.star+n1
    theta1 = (1-2*n.star)/(1+2*n.star);
29
    theta2 = theta1;
    \operatorname{cat}("a=", a, "n1=", n1, "n2=", n2, "n.star=", n.star]
      "theta.pich=", (1-2*n.star)/(1+2*n.star), "\n");
33
    theo.stat.point = c(s.star, i.star, 0.5);
35
    #init
37
    s = 0.8;
    i = 0.2;
39
    r = 0;
    x = 0.20;
41
    state = c(s, i, x); # no r-component
43
45
    # rhs of ODE
    rhs <- function(state){</pre>
47
      s = state[1]; i = state[2];
49
      x = state[3];
      s1 = -bbeta * s * i + (1-x) * B - B * s;
      i1 = bbeta*s*i - alpha*i-B*i;
51
```

```
x1 = -c * x * theta2 * (1-x+n2+a*(1-i)) /
         ((x+n1+a*i)+theta2*(1-x+n2+a*(1-i)));
       x1 = x1+c*(1-x)*theta1*(x+n1+a*i) /
         (\text{theta1}*(x+n1+a*i)+(1-x+n2+a*(1-i)));
       return(c(s1,i1,x1));
57
     }
     alu <- function(x, i){
       return(
61
         -x*theta2*(1-x+n2+a*(1-i)) /
          ((x+n1+a*i)+theta2*(1-x+n2+a*(1-i)))
63
         +(1-x)*theta1*(x+n1+a*i) /
          (\text{theta1}*(x+n1+a*i)+(1-x+n2+a*(1-i)))
65
       );
     }
67
     curve(alu(x, i.star), from=0, to=1);
69
     curve(alu(x, i.star+0.01), from=0, to=1, add=TRUE, col="blue");
     curve(alu(x, i.star+0.1), from=0, to=1, add=TRUE, col="blue");
71
     curve(alu(x, i.star - 0.01), from=0, to=1, add=TRUE, col="red");
     curve(alu(x, i.star - 0.1), from=0, to=1, add=TRUE, col="red");
73
     abline (h=0);
     simul.plot<-function(horizont){</pre>
       # simulate
77
       res <<- numeric();
       aver. stat = c(0, 0, 0); no. aver = 0;
79
       h = 0.001; tt = 0; h = 0.01;
       while (tt<horizont){</pre>
81
         state <<- state+h*rhs(state);</pre>
         t t
                = tt + h;
83
                \ll rbind(res, c(tt, state));
         res
         if (tt>horizont/2){
85
           aver.stat = aver.stat + state;
           no.aver = no.aver + 1;
87
         }
       }
89
       plot(res[,1], res[,2], t="l", ylim=c(0,1),
91
           main = paste("c", as.character(c)),
            ylab="S(black)10*I(blue)x(orange)");
93
       lines(res[,1], 10*res[,3], t="1", col="blue")
95
       lines (res [,1], res [,4], t="l", col="orange")
97
       state <<- state;</pre>
       aver.state <<- aver.stat/no.aver;</pre>
99
     }
101
    my.tol = 1e-6;
103
    \max.iter = 500000;
105
```

```
\# simulate to find a stat state
107
     get.stat.state <- function(init.state){</pre>
       # forward simulation until max iter,
       # or || rhs || <- my.tol
       my.tol2 = my.tol**2;
113
       my.state = init.state;
       h
                 = 0.01;
115
       for (i in 1:max.iter){
         loc.rhs = rhs(my.state);
117
         if (\operatorname{sum}(\operatorname{loc.rhs} * * 2) < \operatorname{my.tol2})
           return (my. state);
119
         }
         my.state = my.state+h*loc.rhs;
121
       }
       cat("# get.stat.state DID NOT CONVERGE!!!\n");
123
       \operatorname{cat}("\#", \operatorname{loc.rhs}, "\backslash n");
       cat("# get.stat.state DID NOT CONVERGE!!!\n");
125
       return (my. state);
     }
     get.jacobian <- function(state){</pre>
131
       # compute jacobian of the vector fieeld at point "state"
             ((f_1)_x, (f_2)_y, (f_3)_z)
       #
       \# J = ( (f_2)_x, (f_2)_y, (f_3)_z )
       #
              ((f_3)_x, (f_2)_y, (f_3)_z)
135
       J = matrix(NA, 3, 3);
       hh = 1e - 3;
137
       rhsp = rhs(state+c(hh, 0, 0));
139
       rhsm = rhs(state+c(-hh,0,0));
       grad = (rhsp-rhsm)/(2*hh);
141
       J[1,] = grad;
       rhsp = rhs(state+c(0, hh, 0));
143
       rhsm = rhs(state+c(0, -hh, 0));
       grad = (rhsp-rhsm)/(2*hh);
145
       J[2,] = grad;
       rhsp = rhs(state+c(0,0,hh));
147
       rhsm = rhs(state+c(0,0,-hh));
       grad = (rhsp-rhsm)/(2*hh);
149
       J[3,] = grad;
       return(J);
     }
155
    # go
157
    # simulate
     if (1==1){
                   # change to (1==1) to enable
161
```

```
nnn = 20;
       c.list = 1+49*(0:nnn)/nnn;
163
              = c.list[1];
       с
       all.res = c();
165
       for (i \quad in \quad 1:(nnn+1)) {
         c = c.list[i];
167
         state = state+c(0.0, 0.01, 0.0)
169
         \operatorname{simul.plot}(300);
171
                = get.jacobian(aver.state);
         J
                = eigen(J);
         ev
173
         all.res = rbind(all.res,
175
         c(c, aver.state, rhs(aver.state), ev$values));
       }
177
179
     }
181
    #generate plot
183
     if (1==1){
185
       11.x = c(); 11.y = c();
       12.x = c(); 12.y = c();
187
       13.x = c(); 13.y = c();
       for (i in 1:(nnn+1)){
189
         11.x = c(11.x, Re(all.res[i,8]));
         l1.y = c(l1.y, Im(all.res[i,8]));
191
         12.x = c(12.x, Re(all.res[i,9]));
         12.y = c(12.y, Im(all.res[i,9]));
193
         13.x = c(13.x, Re(all.res[i,10]));
         13.y = c(13.y, Im(all.res[i,10]));
195
       }
       l.x = c(l1.x, l2.x, l3.x);
197
       l.y = c(l1.y, l2.y, l3.y);
199
       plot(l1.x, l1.y, xlim=c(-0.05, 0.05), ylim=c(min(l.y)),
         max(l.y)), xlab="Re(lambda)", ylab="Im(lambda)");
201
       points(12.x, 12.y, col="blue");
       points(13.x, 13.y, col="green");
203
       abline(h=0, lty=3); abline(v=0, lty=3);
     }
205
```

6.2 R Scripts used for the Data Fitting and Figure 6

6.2.1 analysis_combined_model.R

```
#
   # Estimate the parameters for the combined model.
   #
      (a) For each county, estimate
   #
4
         - Zealot model (beta-distrib)
   #
         - All reinforcement parameters equal
   #
6
         - All parameters can be chosen independently.
   #
   \# (b) Compare the models by the log-likelihood-ration test
8
   \# (c) Compare empirical and theoretical distributions
   #
         by the Kolmogorov-Smirnov tests.
   #
12
    post <- function(nme){</pre>
     \# remove blanks
14
     xx = strsplit(nme, "");
     nme1 = paste(xx[[1]], collapse="");
16
     \operatorname{cat}(\operatorname{nmel}, "\backslash n");
     postscript(file=nme1, pointsize=32, paper="special",
18
     width=8, height=9, horizontal=FALSE);
   }
20
22
   \# read data
   #
24
   vaccination_data = read.csv2('my_path_to_data/vaccination_data.csv',
26
     header=TRUE)
28
   \# read tools
30
   #
   32
    source("parameter_estimation_reinforcement.R");
34
   36
   # initialisation (apply changes ONLY here)
   38
   #initial parameters
40
   theta_hut.init = 0.5;
    psi_hut.init = 0.5;
42
    ksi_hut.init = 0.5;
   s_hut.init
               = 100;
44
46
   #switches for work todo
   analyse = TRUE;
48
   produce.table = FALSE;
   produce.figures = FALSE;
50
```

```
52
   \# prepare the data
54
   myDataEsti = c();
58
    impfer = vaccination_data Wert[1:400]; #last few ones are NA
    myDataEsti = impfer/100;
60
   \#remove all values below 0.5 as we esteem them to be fixed voters
   myDataEsti=myDataEsti[myDataEsti>0.5];
62
   #renormalise it, such that 0.5 \rightarrow 0 and 1 \rightarrow 1
64
   myDataEsti = myDataEsti * 2 - 1;
66
    incidences = vaccination_data $Inzidenz [1:400];
   inc_no_NA = get_rid_of(incidences);
68
   #get the mean value of all incidences w/o the NA ones
   dummy. i = mean(inc_no_NA);
70
   #replace all NA values in incidences by the dummy.i
72
   incidences = replace (incidences, dummy.i);
74
   #set the incidence that will be used to the mean of all observed values
76
   i.param = mean(incidences);
78
   80
   # data fit
   82
84
    if (analyse){
     # check optima
86
     s_hut.max=2500;
88
     res.tab = c();
90
     {
92
       mmean = mean(myDataEsti);
94
       hist (myDataEsti, freq = FALSE, nclass=30,
96
       main="full reinforcement",
       xlim=c(0,1));
98
100
       # first run: estimate the full reinforcement model
       \operatorname{cat}("\operatorname{first} \operatorname{run}", "\backslash n");
       #para.ref in following order: nu_hut, theta_hut, psi_hut, ksi_hut, s_hut
106
```

```
para.ref = c(mean(myDataEsti), theta_hut.init, psi_hut.init,
           ksi_hut.init, s_hut.init); # define init para
108
         lll.last = lll(para.ref);
         cat("lll.last: ", lll.last, "\n");
                                = TRUE;
                                              \# we aim at the full model
         unrestricted.model
         unrestrict.theta_hut
                                = TRUE;
         opti.cyclic(para.ref);
114
         \operatorname{cat}(\operatorname{lll.last}, "\backslash n");
116
         para.unrest = para.last;
                                         \# store the result
         111.unrest = 111.last;
         theta.res
                     = theta_hut*s_hut;
120
        # Kolmogorov-Smirnov test
         res.ks = ks.test(myDataEsti, function(x) { pReinforce(x) });
124
        # second run: reinforcement model, force equal reinforcement
            parameters
        128
         \operatorname{cat}(\operatorname{"second} \operatorname{run"}, \operatorname{"} n");
         para.ref = c(mean(myDataEsti), theta_hut.init, psi_hut.init,
130
           ksi_hut.init, s_hut.init); # define init para
         111.last = 111(para.ref);
         \operatorname{cat}(\operatorname{lll.last}, "\backslash n");
         hist (myDataEsti, freq = FALSE, nclass=30,
134
         main="reinforcement with equal reinf. params",
         xlim=c(0,1));
136
         unrestricted.model
                                = TRUE:
                                              \# we aim at the full model
138
         unrestrict.theta_hut
                                = FALSE;
                                              \# we want to keep equal
            parameters for reinforcement
         opti.cyclic(para.ref);
140
         \operatorname{cat}(\operatorname{lll.last}, "\setminus n");
142
         para.halfRestrict = para.last;
                                                \# store the result
         lll.halfRestrict = lll.last;
         theta.res
                    = theta_hut*s_hut;
144
        # Kolmogorov-Smirnov test
146
         res.halfRestrict.ks = ks.test(myDataEsti, function(x){pReinforce(x)})
            ;
148
        # third run: estimate the zealot model (beta-distrib)
152
        *****
         \operatorname{cat}(\operatorname{"third}\operatorname{run"}, \operatorname{"} n");
154
                                                    # we fix all reinforcement-
         unrestricted.model
                                = FALSE;
            paras
                                = FALSE;
         unrestrict.theta_hut
```

```
theta_hut.zealot = 0;
158
         para.ref = c(mean(myDataEsti), theta_hut.zealot, psi_hut.init,
160
           ksi_hut.init,s_hut.init);
         hist(myDataEsti, freq = FALSE, nclass=30,
162
         main="zealot model",
         xlim=c(0,1));
164
         opti.cyclic(para.ref);
         cat("lll.last: ", lll.last, "\n");
         para.restrict = para.last;
                                                  # store result
168
         111.restrict = 111.last;
         # kolmogorov-smirnov-test
         res.restric.ks = ks.test(myDataEsti, function(x) { pReinforce(x) });
172
         line = c("run", "pro-vaxx",
174
         theta.res,
         para.unrest,
176
         lll.unrest,
         para.halfRestrict,
178
         lll.halfRestrict,
         para.restrict,
180
         lll.restrict,
         pchisq(2*(111.unrest-111.restrict), df=2,
182
         lower.tail=FALSE),
         pchisq(2*(111.unrest-111.halfRestrict), df=2,
184
         lower.tail=FALSE),
186
         res.ks$p.value,
         res.halfRestrict.ks$p.value,
         res.restric.ks$p.value
188
         );
190
         res.tab = rbind(res.tab, line);
       }
       \# names orient themselves ar the supplement II of the paper
       col.names = c(
194
       "run", "opinion"
       "Theta1PlusTheta2.unr",
196
       "nu.unr", "theta.hat.unr", "psi.unr", "ksi.unr", "s.unr",
       "lll.unr",
198
       "nu.halfr"
                  , "theta.hat.halfr", "psi.halfr", "ksi.halfr", "s.halfr",
       "lll.halfr"
200
       "nu.restr"
                   "theta.hat.restr", "psi.restr", "ksi.restr", "s.restr",
       "111.restr"
202
       "lll.unr.rest", "lll.unr.halfr",
       "ks.unres", "ks.halfRestr", "ks.restr");
204
       dimnames (res.tab) [[2]] = col.names;
206
       save(file="datAnaCombinedModel.rSave", res.tab);
208
     }
210
212
```

```
if (produce.table) {
        \# produce a table
214
         load ( file="datAnaMyModel_V1.rSave" );
         sink (file="datCombinedModel.tex");
216
         cat(dimnames(res.tab)[[2]][c(1,2,4,5,6,7,8)]); cat(" theta1 ");
         cat (" theta2 "); cat (dimnames (res.tab) [[2]] [c(22,24,26)]);
218
         \operatorname{cat}("\setminus n");
         nn = dim(res.tab)[1];
         for (j in 1:nn){
           cat (res.tab [j,1], "&", res.tab [j,2], "&");
cat (res.tab [j,4], "&", res.tab [j,5], "&");
cat (res.tab [j,6], "&", res.tab [j,7], "&", res.tab [j,8], "&");
222
           # theta_2 = h.s*h.theta*(1-h.psi)
            cat(as.double(res.tab[j,5])*as.double(res.tab[j,8])*
226
               as.double(res.tab[j,6]), " & ");
            cat(as.double(res.tab[j,5])*as.double(res.tab[j,8])*
            (1-as.double(res.tab[j,6])), "&");
cat(as.double(res.tab[j,22]), "&",
230
            as.double(res.tab[j,24]), "\& ",
            as.double(res.tab[j,26]),
232
            " \setminus \setminus \setminus n";
         }
234
         cat ("Test hat.psi=0.5 versus free model (where is the reinforcement?) \n
236
             ");
         \operatorname{cat}(\operatorname{as.double}(\operatorname{res.tab}[j,23]), " \setminus n");
238
         sink();
240
      }
242
244
      if (produce.figures) {
        # produce figures
246
         impfer = vaccination_data Wert[1:400];
         myDataEsti = impfer/100;
248
         mmean <<- mean(myDataEsti);
250
         post(paste("Combined_model", as.character(j),".eps", sep=""));
         hist(myDataEsti, freq = FALSE,
252
         main=paste("Vaccinational behaviour"),
         xlim=c(0,1), xlab="amount of pro-vaxxers x", nclass=30);
254
         para.last = as.double(res.tab[j, 4:8]);
256
         lll.last = lll(para.last);
258
         cat(lll.last, "\n"); mmean <<< mean(myDataEsti);
260
         \operatorname{curve}(g(x), \operatorname{add}=\operatorname{TRUE}, \operatorname{lwd}=2);
262
         para.last = as.double(res.tab[j, 16:20]);
         111.last = 111(para.last);
264
         \operatorname{cat}(\operatorname{lll.last}, "\backslash n");
         \operatorname{curve}(g(x)), \operatorname{add}=\operatorname{TRUE}, \operatorname{lwd}=2, \operatorname{lty}=2);
266
```

```
dev.off();
268
        post(paste("Combined_model_2", as.character(j),".eps", sep=""));
270
        hist(myDataEsti, freq = FALSE,
        main=paste("Vaccinational behaviour"),
272
        xlim=c(0,1), xlab="amount of pro-vaxxers x", nclass=30);
274
        para.last = as.double(res.tab[j, 4:8]);
276
        111.last = 111(para.last);
        cat(111.last, "\n"); mmean <<- mean(myDataEsti);
278
        curve(g(x), add=TRUE, lwd=2);
280
        para.last = as.double(res.tab[j, 10:14]);
282
        lll.last = lll(para.last);
        \operatorname{cat}(\operatorname{lll.last}, "\setminus n");
284
                                     lwd=2, col = "green");
        curve(g(x), add=TRUE,
286
        para.last = as.double(res.tab[j, 16:20]);
        111.last = 111(para.last);
288
        \operatorname{cat}(\operatorname{lll.last}, "\backslash n");
        \operatorname{curve}(g(x), \operatorname{add}=\operatorname{TRUE}, \operatorname{lwd}=2, \operatorname{lty}=2);
290
        dev.off();
      }
292
```

6.2.2 parameter_estimation_reinforcement.R

```
#
  # Estimate the parameters for the reinforcement model
  #
  \# general functions
6
  #
  8
10
  #
  \# myDataEsti: vector with values in (0,1); the data it uses for the
     estimation
  #
12
   replace <- function(vector, replacement){</pre>
14
    for (i in 1:length(vector)){
      if (is.na(vector[i])){
16
       vector[i] = replacement;
      }
18
    }
20
    return (vector);
   }
22
```

```
get_rid_of <- function(vector){</pre>
       vector_no_NA = c();
24
       last_index = 1;
       for (i in 1:length(vector)) {
26
          if (!is.na(vector[i])){
            vector_no_NA[last_index] = vector[i];
28
            last_index = last_index + 1;
         }
30
       }
       return (vector_no_NA);
32
     }
34
     post <- function(nme){</pre>
       \# remove blanks
36
       xx = strsplit(nme, "");
       nme1 = paste(xx[[1]], collapse="");
38
       \operatorname{cat}(\operatorname{nmel}, "\backslash n");
       postscript(file=nme1, pointsize=28, paper="special",
40
       width=8, height=9, horizontal=FALSE);
     }
42
    44
    #
    # define distribution
46
    #
    48
     s_hut.max = 2000;
50
     f.norm <- function()
52
       # an additive constant to avoid large numbers
       \# in the exponent; this number is chosen in dependence
54
       \# on the mean value of the data at hand.
       x = mmean;
56
       return(-1*(s_hut*theta_hut*(0.5*x**2-phi_hut*x)))
       +\log(x)*((1-\text{theta\_hut})*\text{nu\_hut}*(1-\text{ksi\_hut})+\text{nu\_hut}*\text{ksi\_hut}*i.\text{param})*s\_\text{hut}
58
       +\log(1-x)*((1-\text{theta\_hut})*(1-\text{nu\_hut})*(1-\text{ksi\_hut})+\text{nu\_hut}*\text{ksi\_hut}*(1-\text{i})
           param))*s_hut)
       );
60
     }
62
     f \leftarrow function(x)
       return(
64
       \exp( s_hut*theta_hut*(0.5*x**2-phi_hut*x))
       +\log(x)*((1-\text{theta\_hut})*\text{nu\_hut}*(1-\text{ksi\_hut})+\text{nu\_hut}*\text{ksi\_hut}*i.\text{param})*\text{s\_hut}
66
       +\log(1-x)*((1-\text{theta}_hut)*(1-nu_hut)*(1-ksi_hut)+nu_hut*ksi_hut*(1-i)
           param))*s_hut +f.norm()
68
       )
       );
70
     }
     get.cc <- function(){
72
       \# get the normalisation constant
       CC=(integrate(f, lower=10**-6, upper=1-10**-6) value)**(-1);
74
       return (CC);
```

```
}
76
78
     g<-function(x){
       \# density (note: CC is computed using the constant f.norm().
80
       \# That constant cancels at the end of the day.
        return(
82
       \mathrm{CC}_{*}
        \exp(s_hut*theta_hut*(0.5*x**2-phi_hut*x))
84
       +\log(x)*((1-\text{theta\_hut})*\text{nu\_hut}*(1-\text{ksi\_hut})+\text{nu\_hut}*\text{ksi\_hut}*i.\text{param})*\text{s\_hut}
       +\log(1-x)*((1-\text{theta\_hut})*(1-\text{nu\_hut})*(1-\text{ksi\_hut})+\text{nu\_hut}*\text{ksi\_hut}*(1-\text{i})
86
           param))*s_hut +f.norm()
        )
        );
88
     }
90
     lll.dat <- function(x,nu_hut,theta_hut,phi_hut,ksi_hut,s_hut, CC){
       # log likelihood for one single data point x
92
        return (
        s_hut*theta_hut*(0.5*x**2-phi_hut*x)
94
       +\log(x)*((1-\text{theta\_hut})*\text{nu\_hut}*(1-\text{ksi\_hut})+\text{nu\_hut}*\text{ksi\_hut}*i.\text{param})*\text{s\_hut}
       +\log(1-x)*((1-\text{theta\_hut})*(1-\text{nu\_hut})*(1-\text{ksi\_hut})+\text{nu\_hut}*\text{ksi\_hut}*(1-\text{i}.
96
            param))*s_hut
       +\log(CC)+f.norm()
98
        );
     }
100
     pReinforce.loc \leftarrow function(x) {
       # parameters given by global parameters;
       \# particularly, CC is defined.
        res = integrate(g, lower = 0, upper = x);
104
        return (res$value);
     }
106
     pReinforce <- function(x){
108
        return(sapply(x, pReinforce.loc));
     }
     \# estimate paras:
     # nu_hut, theta_hut, phi_hut, ksi_hut, B
114
     theta_hut.\max = 1800;
     theta_hut.max = 1900;
     tryCatch.W.E <- function(expr){</pre>
118
       W <- NULL
       w.handler \leftarrow function(w) { # warning handler
       W \ll w
        invokeRestart("muffleWarning")
122
        }
        list (value = withCallingHandlers(tryCatch(expr, error = function(e) e),
124
        warning = w. handler),
        warning = W)
126
     }
128
```

```
\# compute log likelihood
     lll <- function(para){</pre>
130
       ppara
                 <-- para;
       nu_hut
                 \ll para [1];
       theta_hut \ll para [2];
                 <\!\!<\!\!- para [3];
134
       phi_hut
       ksi_hut
                 \langle \langle - \text{ para}[4];
       s_hut
                 \ll - \min(s_hut.max, abs(para[5]));
136
      OK = TRUE;
       aa = tryCatch.W.E(get.cc());
138
       aa <\!\!<\!\!- aa;
       if (!(is.double(aa$value))>0) return(-10000);
140
      CC<<-- aa$value;
142
       return (sum(lll.dat(myDataEsti, nu_hut, theta_hut, phi_hut, ksi_hut, s_hut,
          CC)));
144
    }
146
    #
    # for the optimization: vary only one of the parameters
148
    #
    search.p1 \leftarrow function(px){
      # para.last gives the framework; we modify parameter 1 only
      pxx = para.last; pxx[1] = px;
152
       return(lll(pxx));
    }
154
    search.p2 <- function(px){
      \# para.last gives the framework; we modify parameter 2 only
156
       pxx = para.last; pxx[2] = px;
       return( lll(pxx) );
158
     }
    search.p3 \leftarrow function(px){
      \# para.last gives the framework; we modify parameter 3 only
      pxx = para.last; pxx[3] = px;
       return( lll(pxx) );
    }
164
    search.p4 \leftarrow function(px){
      \# para.last gives the framework; we modify parameter 4 only
      pxx = para.last; pxx[4] = px;
       return(111(pxx));
168
    }
172
    #
174
    # optimisation
176
    #
    178
     opti.cyclic <- function(para.init){</pre>
      \# optimise cyclically the parameters.
180
      #
      \# we have different modes
182
```

```
\# unrestricted.model == FALSE: fix all parameters expect of s_hut, and
           nu_hut \implies we can take
                                          the reinforcement to zero and fit a beta
       #
184
            distribution.
         unrestrict.theta_hut
                                   == TRUE:
                                              allow theta_hut to vary.
       #
                                          (if FALSE: we can fix theta_hut=0.5, and
186
       #
            in this way,
       #
                                           try to find out where the reinforcement
            takes place)
188
       mmean <<- mean(myDataEsti);
190
       para.last <- para.init;</pre>
       111.last <- 111(para.ref);
192
       last.s_hut = -1; no.s_hut.const = 0; last.lll=-1e10;
       fertig = FALSE;
194
       i = 0;
196
       while ((i <10000)&(fertig==FALSE)){
         i = i + 1;
198
         cat("para.last:", para.last," \n");
         111.x = 111.1ast;
200
         para.last <<- para.last;</pre>
         res1 = optimize(search.p1, interval=c(0,1), maximum=TRUE);
202
         para.loc1 = para.last; para.loc1[1] = res1 $maximum;
         lll.lok = lll(para.loc1);
204
         if (111.lok>111.last){
            para.last <- para.loc1;</pre>
206
            lll.last <- lll.lok;
         }
208
         ##optimize ksi
210
         para.last <<- para.last;</pre>
         res4 = optimize(search.p4, interval=c(0,1), maximum=TRUE);
212
         para.loc4 = para.last; para.loc4[4] = res4 $maximum;
         111.lok = 111(para.loc4);
214
         if (111.lok>111.last){
216
            para.last <- para.loc4;</pre>
            111.last <- 111.lok;
         }
218
         if (unrestricted.model){
220
            para.last <<- para.last;</pre>
            res2 = optimize(search.p2, interval=c(0,1), maximum=TRUE);
222
            para.loc2 = para.last; para.loc2[2] = res2 $maximum;
            111.lok = 111(para.loc2);
            if (111.lok>111.last)
              para.last <- para.loc2;</pre>
226
              111.last <- 111.lok;
            }
            if (unrestrict.theta_hut) {
230
              para.last <<- para.last;</pre>
              res3 = optimize(search.p3, interval=c(0,1), maximum=TRUE);
232
              para.loc3 = para.last; para.loc3[3] = res3 $maximum;
```

```
111.lok = 111 (para.loc3);
234
               if (lll.lok>lll.last){
                 para.last <- para.loc3;</pre>
236
                 111.last <- 111.lok;
               }
            }
          }
240
          111.1
                     = 111 (para.last);
242
          Delta = 0.01;
          if (para.last[5] > 10)
                                   Delta = 0.1;
244
          if (para.last[5] > 50)
                                   Delta = 0.5;
          if (para.last[5] > 100) Delta=1;
246
          111.p2
                     = 111 (para.last+c(0,0, 0,0, Delta));
248
                     = 111 (para.last+c(0,0,0,0, -Delta));
          111.m2
          if (lll.p2>lll.1) {
250
            paral.loc3 = para.last+c(0,0,0,0, Delta);
            para.last = para.last+c(0, 0, 0, 0, 0, 0);
252
            last.lll = lll.p2;
          else 
254
            if (lll.m2>lll.1) {
               paral.loc3 = para.last+c(0,0,0,0,-Delta);
              para.last <- para.last+c(0,0,0,0,-Delta);
               last.lll
                           <- 111.m2
258
            else 
               paral.loc3 = para.last+c(0,0,0,0,0);
260
               para.last
                          \leftarrow para.last+c(0,0,0,0,0);
               last.lll
                           <- lll.1
262
            }
          }
264
          if (last.s_hut!=para.last[3])
            last.s_hut = para.last[3];
266
            no.s_hut.const = 0;
            111.x = last.111;
268
          else 
            no.s_hut.const = no.s_hut.const+1;
270
            loc.lll = lll(para.last);
            if (last.lll>lll.x+1e-6) {
272
              no.s_hut.const = 0;
            }
274
            if (no.s_hut.const>100) fertig=TRUE;
          }
          para.last <<- para.last;</pre>
278
          lll.last \ll last.lll;
        \operatorname{cat}(i, "", 111(\operatorname{para.last}), "\backslash n");
280
282
        if (fertig) {
          curve(g(x), add=TRUE, col="red");
        }
284
     }
286
   }
```